Machine Learning CS 4641-B Summer 2020



#### Lecture 12. Principle Component Analysis Xin Chen

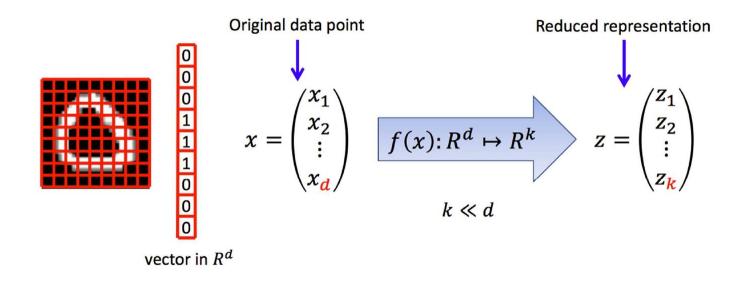
These slides are based on slides from Mahdi Roozbahani

### Outline

- Overview 🗲
- Main idea of Principle Component Analysis (PCA)
- PCA algorithm
- PCA and SVD
- Summary

### What is Dimension reduction?

- The process of reducing the number of features under the consideration:
  - One can combine, transform or select features
  - One can use linear and nonlinear operations



## Applications of the dimension reduction

- The dimension-reduced data can be used for:
  - Visualizing, exploring and understanding the data
  - Aggregating weak signals in the data
  - Cleaning the data
  - Speeding up subsequent learning tasks
  - Building simpler model later
- Key questions of a dimensionality reduction algorithm
  - What is the criterion for carrying out the reduction process?
  - What are the algorithm steps?

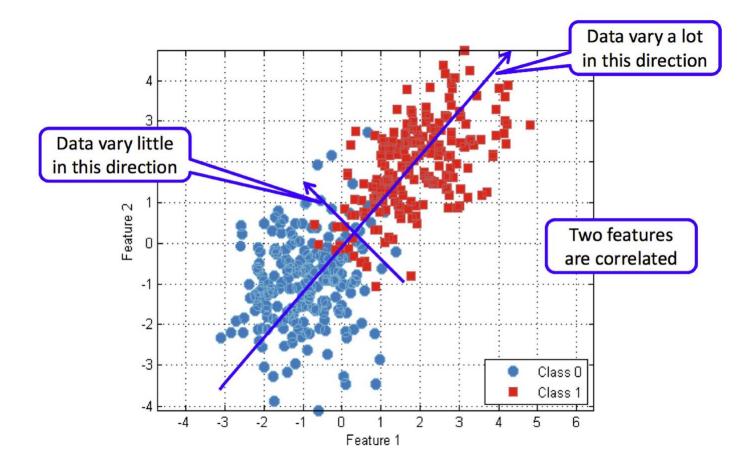
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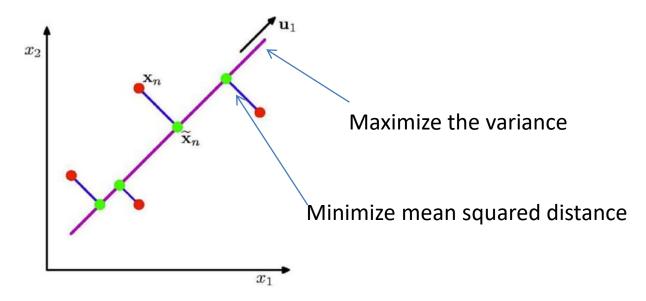
# PCA: Dimension reduction by capturing variation

- There are many criteria, geometric based, information theory based, etc.
- One criterion: want to capture variation in data
  - Variations are "signals" or "useful" information in the data
  - Need to normalize each variable first
- In the process, also discover variables or dimensions highly correlated
  - Represent highly related phenomena
  - Combine them to form a stronger signal
  - Lead to simpler presentation

#### Capture Variation in Data



#### Two perspective of Principal Component Analysis (PCA)



- Orthogonal projection of the data onto a lowerdimension linear space that
  - Maximize variance of projected data
  - Minimize mean squared distance between the data points and projections.

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#### Formulating the problem

Given *n* data points,  $\{x_1, x_2, x_3, ..., x_n\} \in \mathbb{R}^d$ , with their mean  $u = \frac{1}{n} \sum_{i=1}^n x_i$ 

Find a direction 
$$w \in R^d$$
, where  $||w|| = \sqrt{\sum_{j \in d} \omega_j^2} = 1$ 

We constrain the norm of *w* to be equal to 1, to avoid having very large variance in each new dimension.

#### Formulating the problem

Given *n* data points,  $\{x_1, x_2, x_3, ..., x_n\} \in \mathbb{R}^d$ , with their mean  $u = \frac{1}{n} \sum_{i=1}^n x_i$ 

$$|\mathbf{w}|| = \sqrt{\sum_{j \in d} \omega_j^2} = 1$$

Optimization target: the variance of the data along direction w is maximized.  $max \frac{1}{n} \sum_{i=1}^{n} (x_i w - uw)^2$ 

Variance in new feature space.

## Formulate it as an optimization problem

Manipulate the objective with linear algebra

$$\frac{1}{n} \sum_{i=1}^{n} (x_i w - \mu w)^2 = \frac{1}{n} \sum_{i=1}^{n} ((x_i - \mu) w)^2 =$$

$$= \frac{1}{n} \sum_{i=1}^{n} \left( (x_i - \mu) w \right)^T ((x_i - \mu) w) = \frac{1}{n} \sum_{i=1}^{n} w^T (x_i - \mu)^T (x_i - \mu) w$$

$$(AB)^T = B^T A^T$$

$$w^T \left( \frac{1}{n} \sum_{i=1}^{n} (x_i - \mu)^T (x_i - \mu) \right) w = w^T C w$$

Covariance matrix

#### Equivalence to the eigenvalue problem

- Claim max  $w^T C w$
- Form lagrangian function of the optimization problem  $L(w, \lambda) = w^T C w + \lambda (1 w^T w)$
- If w is a maximum of the original optimization problem, then there exists a  $\lambda$ , where  $(w, \lambda)$  is a stationary point of  $L(w, \lambda)$
- This implies that  $\frac{\partial L}{\partial w} = 0 = 2Cw 2\lambda w \Rightarrow Cw = \lambda w$

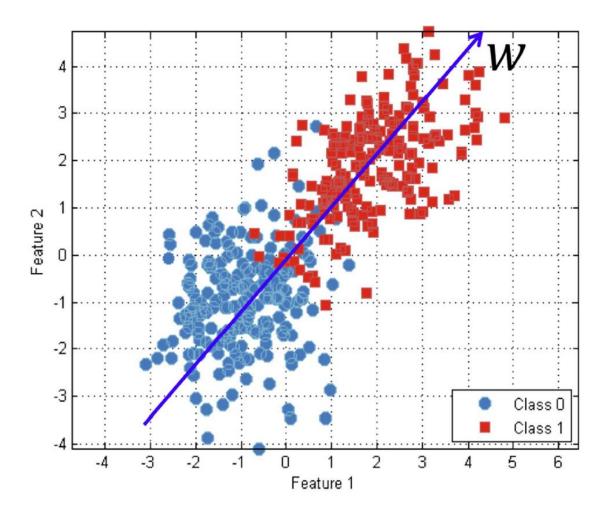
#### Eigen value problem

- Eigen-value problem
  - Given a symmetric matrix  $C \in R^{d \times d}$
  - Find a vector  $w \in \mathbb{R}^d$  and ||w|| = 1
  - Such that  $Cw = \lambda w$

• There will be multiple solutions of the eigenvectors  $w_1, w_2, \dots$  of C corresponding to the largest eigenvalue  $\lambda_1, \lambda_2, \dots, \lambda_d$ 

- They are ortho-normal:  $w_i^T w_i = 1, w_i^T w_j = 0$ 

#### Principle direction of the data



#### Variance in the principle direction

- Principle direction w satisfies  $Cw = \lambda w = w\lambda$
- Variance in principle direction is  $w^T C w = w^T w \lambda = \lambda$ Eigen value

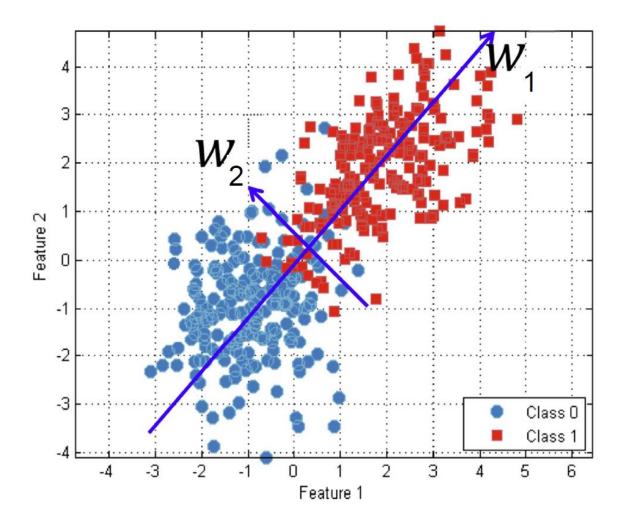
## Multiple principle directions

- Directions  $w_1, w_2, \dots$  which has
  - The largest variances
  - But are orthogonal to each other

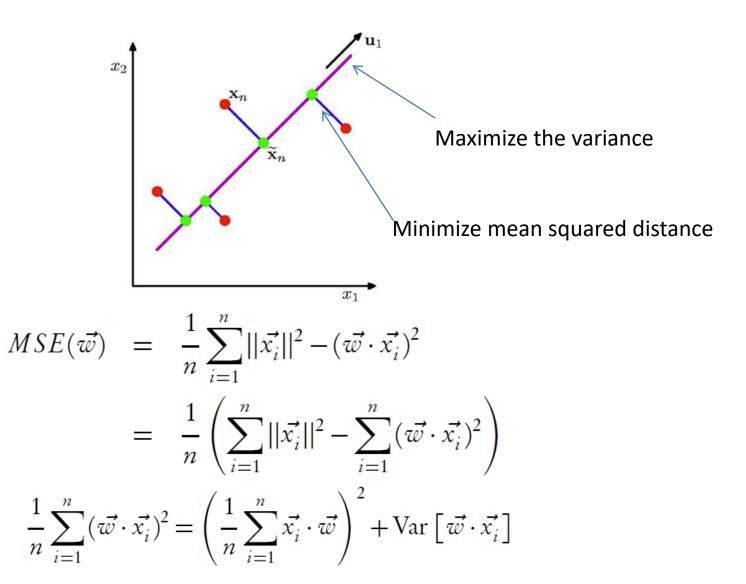
- Take the eigenvectors  $w_1, w_2, \dots$  of C corresponding to
  - The largest eigenvalue  $\lambda_1$
  - The second largest eigenvalue  $\lambda_2$

— ...

#### Extra principle directions



#### Remember the two perspectives



# Relations between principle components

- Principle component #1: points in the direction of largest variance.
- Each subsequent principle component
  - Is orthogonal to the previous ones, and
  - Points in the directions of the largest variance of the residual subspace.

#### The PCA algorithm

Given *n* data points,  $\{x_1, x_2, x_3, ..., x_n\} \in \mathbb{R}^d$ , with their mean  $u = \frac{1}{n} \sum_{i=1}^n x_i$ 

Step 1: Estimate the mean and covariance matrix from data,  $C = \frac{1}{n} \sum_{i=1}^{n} (x_i - u)^T (x_i - u)$ 

Step 2: Take the eigenvectors  $w_1, w_2, ...$  of *C* corresponding to the largest eigenvalue  $\lambda_1$ , the second largest eigenvalue  $\lambda_2, ...$ 

Step 3: Compute reduced representation  

$$z_i = \left(\frac{(x_i - u_1)}{\sigma_1} w_1 \frac{(x_i - u_2)}{\sigma_2} w_2 \dots\right) \qquad \qquad \begin{array}{c} z = n \times k \\ k < d \end{array}$$

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#### **Singular Value Decomposition**

 $X_{n \times d}$  n: instances d: dimensions X is a centered matrix

 $X = U\Sigma V^T$ 

 $U_{n \times n} \rightarrow unitary \ matrix \rightarrow U \times U^T = I$  $\Sigma_{n \times d} \rightarrow diagonal \ matrix$ 

According to PCA  $\rightarrow Cw = \lambda w = w\lambda$ 

Covariance 
$$C_{d \times d} = \frac{1}{n} \sum_{i=1}^{n} (x^i - \mu)^T (x^i - \mu) = \frac{X^T X}{n}$$

$$X = U\Sigma V^{T}$$

$$C = \frac{X^{T}X}{n}$$

$$C = \frac{V\Sigma^{T}U^{T}U\Sigma V^{T}}{n} = \frac{V\Sigma^{2}V^{T}}{n}$$

$$C = \frac{V\Sigma^2 V^T}{n} = V \frac{\Sigma^2}{n} V^T$$

$$CV = V \frac{\Sigma^2}{n} V^T V = V \frac{\Sigma^2}{n}$$
  
According to Eigen-decomposition definition  $\Rightarrow CV = V\Lambda$ 

*V* is the eigen vectors of covariance (Principal directions)

$$\lambda_i = \frac{\sigma_i^2}{n}$$
  $\rightarrow$  The eigenvalues of covariance matrix

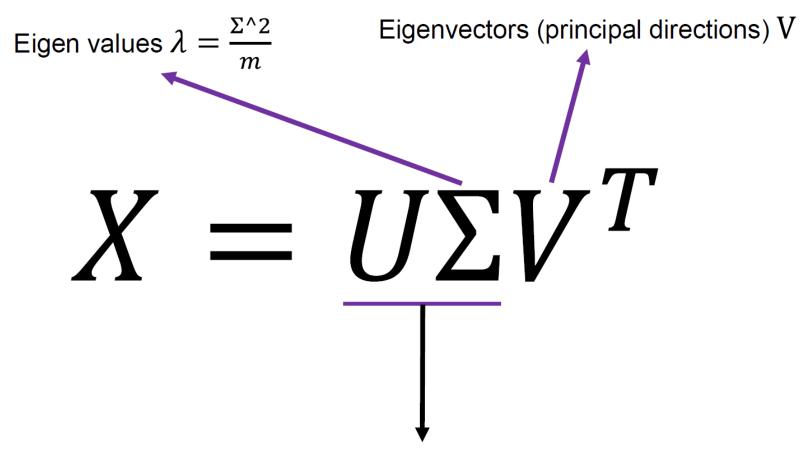
Let's project the data (X) on principal directions:

$$XV = U\Sigma V^T V = U\Sigma$$

XV is independent linear combinations of the original data

Projection of one instance (x) on the first principal direction using k dimensions

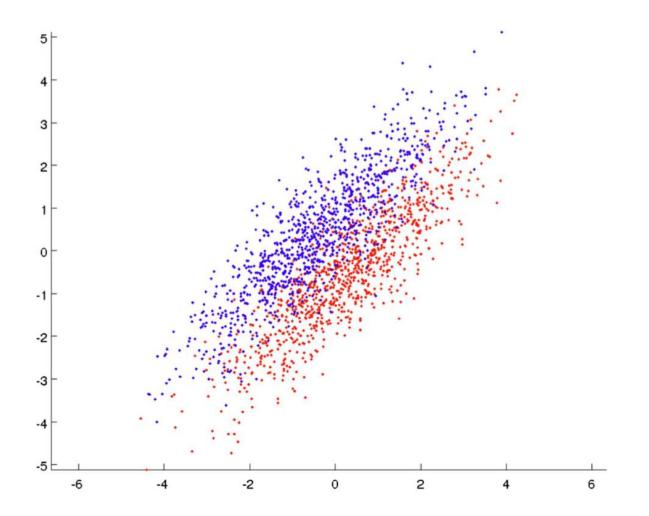
$$\begin{aligned} \mathbf{p}_1 &= \left[ u_{1 \times 1} \Sigma_{1 \times 1} , \, u_{1 \times 2} \Sigma_{2 \times 2} , \, \dots , \, u_{1 \times k} \Sigma_{k \times k} \right] \\ \mathbf{p}_2 &= \left[ u_{2 \times 1} \Sigma_{1 \times 1} , \, u_{2 \times 2} \Sigma_{2 \times 2} , \, \dots , \, u_{2 \times k} \Sigma_{k \times k} \right] \\ & \Sigma \Rightarrow k \times k \\ \text{Upper left corner} \end{aligned}$$



Principal components (Scores) or projections on principal directions

 In fact, using the SVD to perform PCA makes better sense numerically than performing the covariance matrix, since the calculating x<sup>T</sup> x can cause loss of precision.

## Are principal components good for classification?



## Why PCA potentially works in classification?

- The dimension with the largest variance corresponds to the dimension and thus encodes the most information (information theory).
- The smallest eigenvectors often simply represent noise components.